# A lava Attack on the Recovery of Sums of Dense and Sparse Signals

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#### Introduction

- Sparse model :
  - many zeros and a few "large" components.
  - Lasso works well
- Dense model:
  - no large parameters and very many small non-zero parameters
  - Ridge works well

Motivation of this work: sparsity is restrictive in some cases:

- predictions
- nonparametric fitting
- Treatment effect inference with many controls.

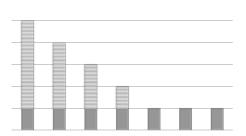
In these applications, variable selection is not a requirement.

# A dense+sparse model

A basic assumption for non-sparse models:

$$\theta = \underbrace{\beta}_{\text{dense signal}} + \underbrace{\delta}_{\text{sparse signal}}.$$

Figure: dense+sparse decomposition



## lava: a new technique for signal recovery

Let  $\ell(\text{data}, \theta)$  be a loss function.

$$\widehat{\theta}_{\text{lava}} = \widehat{\beta} + \widehat{\delta},$$

where

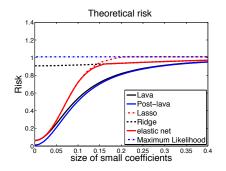
$$(\widehat{\beta}, \widehat{\delta}) = \arg \min_{(\beta', \delta')' \in \mathbb{R}^{2p}} \Big\{ \ell(\text{data}, \beta + \delta) + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\delta\|_1 \Big\}.$$

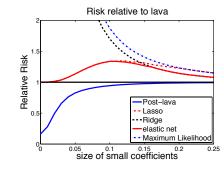
•  $\ell_2$ -part captures dense signal;  $\ell_1$ -part captures sparse signal.

# Risk comparison in $Z \sim N(\theta, I)$

$$\begin{split} \theta &= (3,q,...,q)', \qquad q : \mathsf{small coefficient} \\ \widehat{\theta} &= \widehat{\beta} + \widehat{\delta}, \quad (\widehat{\beta},\widehat{\delta}) = \arg\min \|Z - \beta - \delta\|_2^2 + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\delta\|_1 \end{split}$$

Figure:  $\mathbb{E}\|\widehat{\theta}(Z) - \theta\|_2^2$ , oracle tunings





#### one-dimensional case

Consider shrinkage estimation:

$$d(Z) = \arg\min_{\theta} (Z - \theta)^2 + P_{\lambda}(\theta)$$

We set

$$P_{\lambda}(\theta) = \lambda_2 |\beta|^2 + \lambda_1 |\delta|, \quad \theta = \beta + \delta$$

- To compare with related methods:
  - Lasso:  $P_{\lambda}(\theta) = \lambda |\theta|$
  - elastic net:  $P_{\lambda}(\theta) = \lambda_2 |\theta|^2 + \lambda_1 |\theta|$
  - Ridge:  $P_{\lambda}(\theta) = \lambda |\theta|^2$

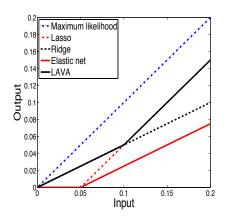
Weighted average of the soft-thresholding and the data.

$$d_{\text{lava}}(Z) = \widehat{\beta} + \widehat{\delta}$$
  
=  $(1 - k)Z + k(\text{soft th.}), \quad k = \frac{\lambda_2}{1 + \lambda_2}$ 

By shrinking towards the data, robust to non-sparse signals.

• Does not produce sparse solutions.

Figure: Shrinkage functions



# lava in the regression model

$$Y = X\theta_0 + U, \quad U \sim N(0, \sigma_u^2 I_n),$$

$$\begin{array}{lcl} \widehat{\theta}_{\mathrm{lava}} & = & \widehat{\beta} + \widehat{\delta}, \\ (\widehat{\beta}, \widehat{\delta}) & = & \arg\min_{\beta, \delta \in \mathbb{R}^p} \frac{1}{n} \|Y - X(\beta + \delta)\|_2^2 + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\delta\|_1. \end{array}$$

# Computations

• If we knew  $\delta$ , then ridge solution :

$$\widehat{\beta}(\delta) = (X'X + n\lambda_2 I_p)^{-1} X'(Y - X\delta).$$

• Substitute  $\beta = \widehat{\beta}(\delta)$  into the objective function,

$$\widehat{\delta} = \arg\min_{\delta \in \mathbb{R}^p} \left\{ \frac{1}{n} \| \mathbf{Y} - \mathbf{X}(\widehat{\beta}(\delta) + \delta) \|_2^2 + \lambda_2 \| \widehat{\beta}(\delta) \|_2^2 + \lambda_1 \| \delta \|_1 \right\}.$$

• So lava is given by:

$$\widehat{\theta} = \widehat{\beta}(\widehat{\delta}) + \widehat{\delta}.$$

# De-densify: another look at Lava

## Theorem (A Key Characterization of the Profiled Lava Program)

Define ridge-projection matrices,

$$P_{\lambda_2} = X(X'X + n\lambda_2I_p)^{-1}X'$$
 and  $K_{\lambda_2} = I_n - P_{\lambda_2}$ ,

and transformed data,  $\widetilde{Y} = K_{\lambda_2}^{1/2} Y$  and  $\widetilde{X} = K_{\lambda_2}^{1/2} X$ . Then

$$\widehat{\delta} = \arg\min_{\delta \in \mathbb{R}^p} \left\{ \frac{1}{n} \|\widetilde{Y} - \widetilde{X}\delta\|_2^2 + \lambda_1 \|\delta\|_1 \right\}.$$

## De-densify: another look at Lava

- In other words, "de-densify" first, then lasso
  - **Step 1:** Ridge-projection matrices,

$$P_{\lambda_2} = X(X'X + n\lambda_2I_p)^{-1}X'$$
 and  $K_{\lambda_2} = I_n - P_{\lambda_2}$ ,

and transformed data,  $\widetilde{Y}=\mathsf{K}_{\lambda_2}^{1/2}\,Y$  and  $\widetilde{X}=\mathsf{K}_{\lambda_2}^{1/2}\,X.$ 

- **Step 2:** Run lasso on  $(\widetilde{Y}, \widetilde{X})$ .
- Why are the signals for the "transformed data" sparse?

$$ilde{Y} = ilde{X}\delta + ilde{U} + \underbrace{ extstyle ilde{V}_{\lambda_2}^{1/2} X eta_0}_{ extstyle extst$$

ullet Taking the transformation  ${\sf K}_{\lambda_2}^{1/2}$  removes the dense component.

## Choices of tuning parameters

#### Data-driven choices

• min-SURE: Suppose  $\widehat{R}(\widehat{\theta}_{\lambda})$  is Stein's Unbiased Risk Estimator for method  $\widehat{\theta}_{\lambda}$ ,

$$\arg\min_{\lambda} \widehat{R}(\widehat{\theta}_{\lambda})$$

K-fold cross validation.

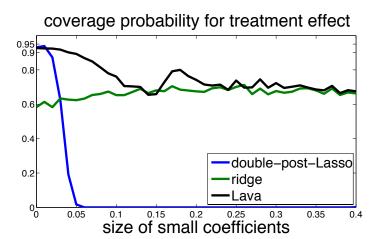
## Inference with many controls

Consider the model

$$y_i = d_i \alpha + X_i' \theta + e_i$$
  
 $d_i = X_i' \gamma + u_i$ 

Belloni et al. (14) used double-post-selection.

- What if  $\theta, \gamma = \text{dense} + \text{sparse}$ ?
- ullet Obtain confidence intervals for lpha that is more robust to the signal
- Example:  $\theta = \gamma = (3, q, ..., q)$ ; where q is the small coefficient.



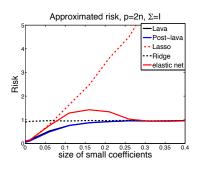
## Monte-Carlo

- n = 100, p = 2n.
- Gaussian regression,

$$\theta = (3, q, ..., q)',$$

- The tuning parameters are selected by numerically minimizing the SURE and 5-fold CV.
- Consider an independent design  $X \sim N(0, I)$ .
- Calculate averaged  $\frac{1}{n} ||X\widehat{\theta} X\theta_0||_2^2$  from 100 replications.

Figure: Risk comparisons: tuning chosen by 5-fold CV



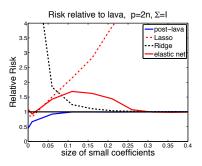
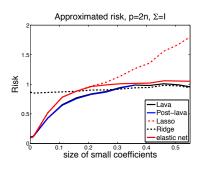
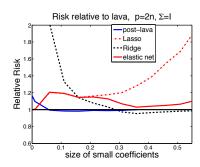


Figure: Risk comparisons: tuning chosen by min-SURE





## Theorem (Deviation Bounds for Lava in Regression)

We have that with probability  $1 - \alpha - \epsilon$  (note that  $\| \mathsf{K}_{\lambda_2} \| \leq 1$ )

$$\begin{split} & \frac{1}{n} \| X \widehat{\theta}_{\mathrm{lava}} - X \theta_0 \|_2^2 \leq \frac{2}{n} \| \operatorname{K}_{\lambda_2}^{1/2} X (\widehat{\delta} - \delta_0) \|_2^2 \| \operatorname{K}_{\lambda_2} \| + \frac{2}{n} \| \operatorname{D}_{\mathrm{ridge}} (\lambda_2) \|_2^2 \\ \leq & \inf_{(\delta_0', \beta_0')' \in \mathbb{R}^{2p}: \delta_0 + \beta_0 = \theta_0} \left\{ \left( B_1(\delta_0) \vee B_2(\beta_0) \right) \| \operatorname{K}_{\lambda_2} \| + \underbrace{B_3 + B_4(\beta_0)}_{\text{bound of } \operatorname{D}_{\mathrm{ridge}}(\lambda_2)} \right\}, \end{split}$$

$$\begin{split} B_1(\delta_0) &= \frac{2^3 \lambda_1^2}{\iota^2(c,\delta_0,\lambda_1,\lambda_2)} \leq \frac{2^5 \sigma_u^2 c^2 \bar{V}_{\lambda_2}^2 \log(2p/\alpha)}{n \iota^2(c,\delta_0,\lambda_1,\lambda_2)}, \\ B_2(\beta_0) &= \frac{2^5}{n} \| \, \mathsf{K}_{\lambda_2}^{1/2} \, X \beta_0 \|_2^2 = 2^5 \lambda_2 \beta_0' \, S(S+\lambda_2 I)^{-1} \beta_0, \\ B_3 &= \frac{2^2 \sigma_u^2}{n} \left[ \sqrt{\mathsf{tr}(\mathsf{P}_{\lambda_2}^2)} + \sqrt{2} \sqrt{\|\,\mathsf{P}_{\lambda_2}^2\|} \| \sqrt{\log(1/\epsilon)} \right]^2, \\ B_4(\beta_0) &= \frac{2^2}{n} \| \, \mathsf{K}_{\lambda_2} \, X \beta_0 \|_2^2 = 2^2 \beta_0' \, V_{\lambda_2} \beta_0 \leq 2^3 B_2(\beta_0) \| \, \mathsf{K}_{\lambda_2} \, \|. \end{split}$$

### Remarks

- **1** Does not require identification of  $(\beta_0, \delta_0)$ . "inf" finds the best split.
- 2 In dense models, lava works similarly to ridge.
- In sparse models, lava works similarly to lasso.

## Conclusions

- Lava is designed for "sparse+dense" models.
- Complements other approaches to structured sparsity: fused sparsity, matrix decomposition, etc.
- Extendable to more general M- and Z- estimations.