

The Case Against JIVE

Davidson and McKinnon (2004) JAE 21: 827-833

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Instrumental variables performance with weak instruments

- Davidson and McKinnon (2004). The case Against JIVE. *Journal of Applied Econometrics* 21: 827-833.
- Blomquist and Dahlberg (1999). Small sample properties of LIML and Jackknife IV estimators: Experiments with Weak Instruments. *Journal of Applied Econometrics* 14: 69-88.
- Angrist, J., W. Imbens and A. Krueger (1999). Jackknife Instrumental Variables Estimation. *Journal of Applied Econometrics* 14: 57-67.

Summarizing

Finite-sample properties Evaluation (According to each study)

	UJIVE	LIML	Comment
DM		✓	LIML the best in reducing bias
BD	?	?	Hard to find a winner!
AIK	✓		A reduced space of parameters

The System of Equations

We have the following system of equations (in matricial terms);

$$Y = X\beta + \varepsilon \quad (1)$$

$$X = Z\pi + \eta \quad (2)$$

where X , Z and η are matrices of dimension $n \times L$, $n \times k$ and $n \times L$ respectively. The number of overidentified restrictions can be calculated as $r = k - L$. Also, there are M common elements in X and Z , then M columns of $n \times L$ matrix η are zero. The endogeneity comes from the following expression

$$E[\varepsilon_i \eta'_i | Z] = \sigma_{\varepsilon\eta} \quad (3)$$

UJIVE 1

JIVE removes the dependence of the constructed instrument $Z_i \hat{\pi}$ on the endogenous regressor for observation i by using the following estimator

$$\tilde{\pi}(i) = (\mathbf{Z}(i)' \mathbf{Z}(i))^{-1} \mathbf{Z}(i)' \mathbf{X}(i) \quad (4)$$

the estimate of the optimal instrument is $Z_i \tilde{\pi}(i)$; then, because ε_i is independent of X_j if $j \neq i$ we claim that

$$E[\varepsilon_i Z_i \tilde{\pi}(i)] = 0$$

this is easily verifiable

$$\begin{aligned} E[E[\varepsilon_i Z_i' \tilde{\pi}(i)] | Z] &\equiv E\left[Z_i (\mathbf{Z}(i)' \mathbf{Z}(i))^{-1} \mathbf{Z}(i)' E[\varepsilon_i \mathbf{X}(i)] \middle| Z\right] \\ &= 0 \end{aligned}$$

See Phillips and Hale (1977) for details.

UJIVE 1

Thus, $\hat{X}_{i,UJIVE} = Z_i \tilde{\pi}(i)$; then the estimator of β is

$$\hat{\beta}_{UJIVE} = \left(\hat{\mathbf{X}}'_{UJIVE} \mathbf{X} \right)^{-1} \hat{\mathbf{X}}'_{UJIVE} \mathbf{Y} \quad (\text{UJIVE 1})$$

We require to perform the estimator in (4) by each observation i .

[AIK] show a *sort of* shortcut

$$Z_i \tilde{\pi}(i) = \frac{Z_i \hat{\pi} - h_i X_i}{1 - h_i}$$

where $h_i = Z_i (Z'Z)^{-1} Z_i'$

Limited Information Maximum Likelihood

We can estimate the parameters in linear regressions with endogenous regressors by using the limited-information-maximum-likelihood estimator. The likelihood is based on normality for the reduced form errors and with covariance matrix, although consistency and asymptotic normality of the estimator do not rely on this assumption. The log likelihood function is

$$L = \sum_{i=1}^N -\ln(2\pi) - \frac{1}{2} \ln |\Omega| - \frac{1}{2} \begin{pmatrix} Y_i - (Z_i\pi)' \beta \\ X_i - Z_i\pi \end{pmatrix}' \Omega^{-1} \begin{pmatrix} Y_i - (Z_i\pi)' \beta \\ X_i - Z_i\pi \end{pmatrix} \quad (\text{LIML}) \quad (5)$$

Experiments

DM have the following system of equations (*in matricial terms*) in order to do the simulations;

$$Y = \iota\beta_1 + x\beta_2 + \varepsilon \quad (\text{Structural eq.})$$

$$x = \sigma_\eta (Z\pi + \eta) \quad (\text{Reduced eq.})$$

where $X = [\iota \ x]$, Z and η are matrices of dimension $n \times 2$, $n \times l$ and $n \times 1$ respectively. The number of overidentified restrictions can be calculated as $r = l - 2$. The elements of ε and η have variances σ_ε^2 and 1 respectively, and correlation ρ .

In order to start with the simulations we need to impose values for parameters. For this, an important guide is the size of the ratio $\|\pi\|^2$ to σ_η^2 .

Setting up parameters

- DM fix values of the π_j to be equal excepting $\pi_1 = 0$.
- The parameter which does varies is denoted by $R_\infty^2 = \frac{\|\pi\|^2}{\|\pi\|^2 + \sigma_\eta^2}$
(This is the asymptotic R^2)
- R_∞^2 is monotonically increasing function of the the ratio $\|\pi\|^2$ to σ_η^2 .
- A small value of R_∞^2 implies that the instruments are weak.
- In the experiments, DM vary the sample size n , the number of overidentifying restrictions (r), the correlation between errors (ρ) and R_∞^2

Performance evaluation

- Because LIML and JIVE estimators have no moments (see Davison and Mackinnon; 2007). DM reports the median bias, this is

$$\text{median bias} = \beta_{d0.5} - \beta$$

- As measure of dispersion DM reports the nine decile range, this is

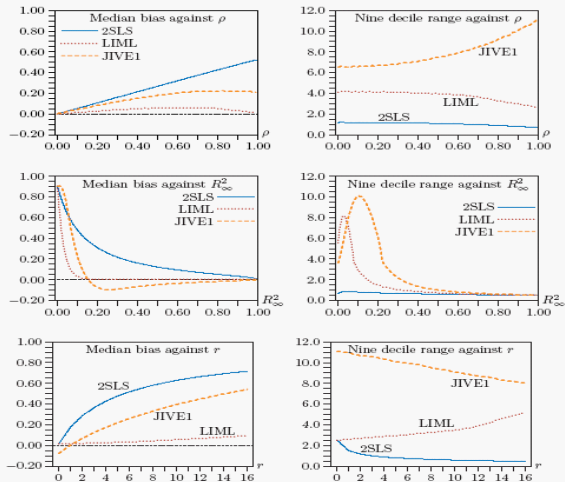
$$9d - \text{range} = \beta_{d0.95} - \beta_{d0.05}$$

- Ackberg and Devereux (2006) suggest to consider

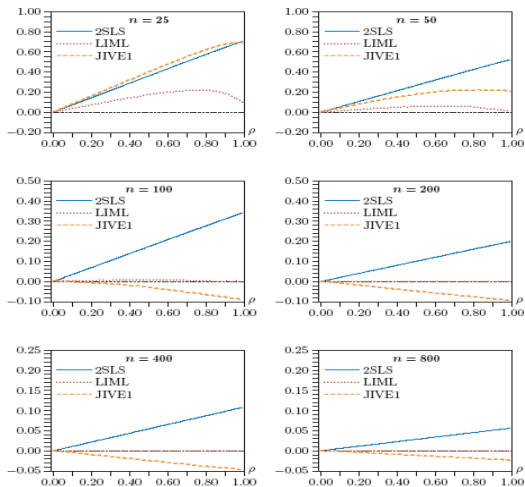
$$\text{Trimmed mean bias} = \frac{1}{n} \sum_j \beta_j - \beta$$

where $\beta_j \in [\beta_{d0.99} \beta_{d0.01}]$.

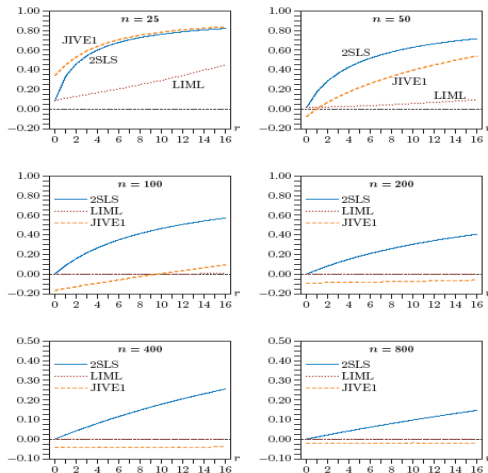
Results (I): Median bias evaluation

Figure A. Median bias and nine decile range when $n = 50$

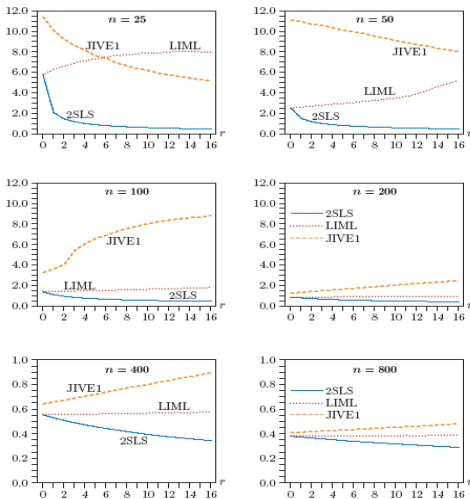
Results (II): Median bias evaluation

Figure 1. Median bias of three estimators, $r = 5$, $R^2_\infty = 0.1$

Results (III): Median bias evaluation

Figure 3. Median bias of three estimators, $R^2_{\infty} = 0.1$, $\rho = 0.9$

Results (V): Dispersion

Figure 6. Nine decile range of three estimators, $R_{20}^2 = 0.1$, $\tau = 0.9$

Conclusions

- 1 AIK and BD give divergent views of JIVE performance. DM's paper re-examine this issue.
- 2 DM conclude that in most regions of the parameters space they have studied, JIVE is inferior to LIML regarding to median bias, dispersion and reliability of inference.
- 3 LIML should be preferred whenever you need to deal with estimators which have no moments.
- 4 DM points out that, however, montecarlo simulations does no support unambiguously the usage of LIML; when the instruments are weak the dispersion is significant in this estimator.

References

-  Angrist, J., W. Imbens and A. Krueger (1999). Jaccknife Instrumental Variables Estimation. *Journal of Applied Econometrics* 14: 57-67.
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-  Davidson, R., and J. McKinnon (2004). The case Against JIVE. *Journal of Applied Econometrics* 21: 827-833.
-  Davidson, R., and J. McKinnon (2006). Reply to Akerberg and Devereux and Blomquist and Dhalberg on "The case Against JIVE". *Journal of Applied Econometrics* 21: 843-844.